

## EXCHANGE RATE VOLATILITY AND EXPORT GROWTH: EVIDENCE FROM SELECTED SOUTH ASIAN COUNTRIES

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### Abstract

The exchange rate regime and the related issues are one of the important yardsticks of the macroeconomic management in striving for economic development through improving the performance of foreign sector. The purpose of this study is to investigate the impact of exchange rate volatility on exports of three South Asian countries, India, Pakistan and Sri Lanka. Using cointegration and vector error correction model (VECM) techniques for the period 1960 to 2007, our empirical findings indicate the presence of a unique cointegrating vector linking real exports, relative export prices, foreign economic activity and real exchange rate volatility in the long run. Real exchange rate volatility exerts significant negative effects on exports both in the short run and the long run. Our results also reveal that improvements in the terms of trade (represented by declines in the real exchange rate) and real foreign income exert positive effects on export activity. Overall, our findings suggest that exporting activities of these South Asian countries can be further boosted up by policies aimed at achieving and maintaining a stable competitive real exchange rate. JEL Classifications: C32, F14, F31.

**Keywords:** Exchange rate volatility, Exports, GARCH, Cointegration, Vector error correction.

### 1. Introduction

The exchange rate regime and the related issues are one of the important yardsticks of the macroeconomic management in striving for the economic development through consistent improvements in foreign sector of the economy. Exchange rates across the world have fluctuated widely particularly after the collapse of the Bretton Woods system of fixed exchange rates. Since then, there has been extensive debate about the impact of exchange rate volatility on international trade. The choice of regime can affect economic growth directly through its effects on the adjustment of the economy to economic shocks, and indirectly through its impact on other important determinants of growth, such

as international trade, investment, capital flows, financial sector and monetary institution development. But the most important determinant is export that is highly affected by the fluctuation of exchange rate. The ability to export goods helps an economy to grow by selling more overall goods and services. One of the core functions of diplomacy and foreign policy within governments is to foster trade in ways that benefit both parties involved.

Exchange rate volatility refers to the amount of uncertainty or risk about the size of changes in a currency's value. A higher volatility means that a currency's value can potentially be spread out over a larger range of values. This means that the price of the currency can change dramatically over a short time period in either direction. A lower volatility means that a currency's value does not fluctuate drastically, but changes in value at a steady pace over some time period. The most commonly held belief is that greater exchange rate volatility generates uncertainty thereby increasing the level of riskiness of trading activity and this will eventually depress trade. It is important to point out here that in less developed countries (LDCs) where the forward markets are less developed and the cost of adjusting to changes in the economic environment is higher, exchange rate volatility coupled with protectionism, could have a major impact on trade and income. It is well documented in the literature that one of the major shortcomings of developing countries is the absence of or underdeveloped financial markets. Because of this, developing economies incur higher transactions costs. Intuitively, one would conclude that it is in the developing countries that the need to understand the policy implications of volatile exchange rates is paramount (*Onafowora and Owoye, 2007*). By using the forward markets and by managing the timing of payments and receipts, a firm can reduce the uncertainties in the short run. The aim of this paper is to provide a contribution to the empirical debate on the relationship between exchange rate volatility and exports for India, Pakistan and Sri Lanka.

Rest of the study is organized as follows. Section 2 contains the survey of some important past studies in order to give the theoretical and empirical evidence in multidimensional support to the topic. Section 3 presents the empirical model and the estimation technique. Section 4 provides empirical results and their interpretation along with nature and sources of data. Final section concludes the study with some policy implications based on the empirical findings of the study.

## 2. Literature Review

There exists an abundance of studies on this topic that have been undertaken internationally, both at theoretical and empirical levels. Two most popular and related approaches have been used in the analysis of trade and exchange rate volatility. One approach is to estimate a simple export demand equation generally with real exports as dependent variable and exchange rate volatility together with relative prices and a measure of economic activity variable as regressors. The other approach is to use the so-called gravity equation models, which explain bilateral trade flows between countries as depending positively on the product of their GDPs and negatively on their geographical distance from each other.

Although the empirical research does not provide a definite result that increased uncertainty has reduced exports, the majority of these studies have found that there is a negative relationship between exchange rate volatility and export performance. The empirical studies that have found a negative relationship between exchange rate volatility and exports include *Thursby and Thursby (1985)*, *Kenen and Rodrik (1986)*, *Koray and Lastrapes (1989)*, *Kumar and Dhawan (1991)*, *Pritchett (1991)*, *Pozo (1992)*, *Savvides (1992)*, *Chowdhury (1993)*, *Arize (1995)*, *Dell' Ariccia (1998)*, *Virgil (2000)*, *Doganlar (2002)*, *Esquivel and Felipe (2002)* and *Onafowora and Owoye (2007)*. On the other hand, *Asseery and Peel (1991)* and *Todani and Munyama (2005)* have reported a positive relationship between exchange rate volatility and exports. Whereas, *Gotur (1985)*, *Solakoglu (1998)*, *De Vita and Abbott (2004)*, *Hondroyannis et al. (2006)*, *Rey (2006)*, and *Boug and Andreas (2007)* could not find any significant relationship between these two time series. The conclusion drawn from empirical literature is that earlier studies tended to find insignificant relationship between export and exchange rate volatility. Cases where significant relations were found, it was both negative and positive.

## 3. Model, Data and Estimation Technique

At the theoretical level, the effects of a greater volatility of exchange rates on export flows are much debated. The literature gives results, which contrast strongly. Using a traditional export demand model with an addition of a measure of exchange rate volatility, the long run export demand function can be written as:

$$RX_t = q_0 + q_1 RP_t + q_2 VOL_t + q_3 Y_t + v_t \quad (1)$$

where,  $RX_t$  is real exports (nominal export/export price index),  $RP_t$  is relative prices (home export price index/industrial countries' export price index). The-

oretically, the bilateral relative price variable should be the ratio of an index of export prices, for the exporting country, and an index of prices of similar goods in the importing country, expressed in the same currency. Since such a measure is not available, in this article the relative price variable (*RP*) is proxied by the real exchange rate.  $VOL_t$  is the real exchange rate volatility which measures uncertainty associated with fluctuations in the exchange rate.  $Y_t$  is the foreign economic activity (industrial production in industrialized countries) which is an indicator of exports of selected South Asian countries and the subscript  $t$  symbolizes the time. Theoretical priors dictate that we should expect  $\rho_1 > 0$  and  $\rho_3 > 0$  and as discussed in the section 2, the sign of  $\rho_2$  is theoretically ambiguous.

The study has used *RX*, *RP* and *Y* in natural logarithm for carrying out the empirical exercise. In order to ensure consistency in data, the exports of each country have been measured in US Dollar. All the data have been collected from *International Financial Statistics*, IMF. The set of data consists of yearly data and spans the period from 1970 to 2007.

Since *Engle (1982)*, the exchange rate volatility has essentially been defined by ARCH (*Autoregressive Conditional Heteroskedasticity*) models, and subsequent generalizations (GARCH, IGARCH, etc.). As *Baillie and McMahon (1989)* and others show, ARCH type effects remain very strong in high-frequency data, but diminish with monthly or quarterly series. We have constructed a GARCH measure of volatility as follows:

$$RER_t = \gamma_0 + \gamma_1 RER_{t-1} + u_t \quad (2)$$

$$\sigma_t = \alpha_0 + \alpha_1 u_{t-1}^2 + \alpha_2 \sigma_{t-1} \quad (3)$$

where  $RER_t$  is real exchange rate expressed in natural logarithm and  $u_t$  is a random error. The conditional variance equation in (3) is a function of three terms: (i) the mean,  $\alpha_0$ ; (ii) news about volatility from the previous period, measured as the lag of the squared residual from the mean equation,  $u_{t-1}^2$  - (the ARCH term); and (iii) the last period's forecast error variance,  $\sigma_{t-1}$  - (the GARCH term).

### 3.1. Unit root test

Since macroeconomic time-series data are usually non-stationary (*Nelson and Plosser, 1982*) and thus conducive to spurious regression, we test for stationarity of a time series at the outset of cointegration analysis. For this purpose, we conduct an augmented Dickey-Fuller (ADF) test, which is based on the t-ratio of the parameter in the following regression.

$$\Delta X_t = \kappa + \phi t + \Theta X_{t-1} + \sum_{i=1}^n \varphi_i \Delta X_{t-i} + \varepsilon_t \quad (4)$$

where  $X$  is the variable under consideration,  $\Delta$  is the first difference operator,  $t$  captures any time trend,  $\varepsilon_t$  is a random error, and  $n$  is the maximum lag length. The optimal lag length is identified so as to ensure that the error term is white noise. While  $\kappa$ ,  $\varphi$ ,  $\Theta$  and  $\phi$  are the parameters to be estimated. If we cannot reject the null hypothesis  $\Theta=0$ , then we conclude that the series under consideration has a unit root and is therefore non-stationary.

### 3.2. Cointegration Test

The econometric framework used for analysis in the study is the *Johansen (1998)* and *Johansen and Juselius (1990)* Maximum-Likelihood cointegration technique, which tests both the existence and the number of cointegration vectors. This multivariate cointegration test can be expressed as:

$$Z_t = A_1 \Delta Z_{t-1} + A_2 \Delta Z_{t-2} + \dots + A_k \Delta Z_{t-k} + u_t \quad (5)$$

where  $Z_t = (RX_t, P_t, Y_t, VOL_t)$ . It can be reformulated in a vector error correction model (VECM) as follows:

$$\Delta Z_t = \Gamma_1 \Delta Z_{t-1} + \Gamma_2 \Delta Z_{t-2} + \dots + \Gamma_{k-1} \Delta Z_{t-k+1} + \Pi Z_{t-1} + u_t \quad (6)$$

where  $\Gamma_i = (I - A_1 - A_2 - \dots - A_k)$  ( $i = 1, 2, 3, \dots, k-1$ ) and  $\Pi = -(I - A_1 - A_2 - A_3 - \dots - A_k)$ . The  $\Pi$  matrix contains information regarding the long run relationships.  $\Pi$  can be factored into  $\alpha\beta'$  where  $\alpha$  will include the speed of adjustment to the equilibrium coefficients while the  $\beta$  will be the long run matrix of coefficients. To determine the number of cointegrating vectors, Johansen developed two likelihood ratio tests: Trace test ( $\lambda_{\text{trace}}$ ) and maximum eigenvalue test ( $\lambda_{\text{max}}$ ). If there is any divergence of results between these two tests, it is advisable to rely on the evidence based on the  $\lambda_{\text{max}}$  test because it is more reliable in small samples (see *Dutta and Ahmed, 1997* and *Odhiambo, 2005*).

## 4. Estimation and Interpretation of Results

The first step in cointegration analysis is to test the unit roots in each variable<sup>1</sup>. Table 1 reports the results of the ADF tests for the level as well as for the

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1. Since the cointegration methodology involves finding a stationary linear combination of a set of variables, which are themselves non-stationary, therefore, a precondition for cointegration to be held is that all variables should be non-stationary.

first-difference of the relevant variables. It is evident that all variables are non-stationary in their levels. However, they are stationary in their first differences. This implies that all the series are integrated of order one [i.e.  $I(1)$ ].

**TABLE 1**  
Augmented Dickey Fuller (ADF) Unit Root Tests

	Export		Relative Price		Volatility		Foreign Economic Activity	
	I(0)	I(1)	I(0)	I(1)	I(0)	I(1)	I(0)	I(1)
India	2.11	-7.33***	1.22	-4.02***	-1.07	-5.35***	-1.57	-4.22***
Pakistan	-0.46	-4.12***	-0.67	-5.53***	-1.56	-5.21***	-	-
Sri Lanka	-0.78	-6.66***	-1.30	-6.58***	-1.72	-6.18***	-	-

\*\*\*Significant at 1% level.

In the next step, we have determined the optimal lag length because Johansen method is known to be sensitive to the lag length. As for as this study is concerned, the Schwarz Bayesian Criteria (SBC) has suggested a lag length of 1 for India, Pakistan and Sri Lanka respectively as optimal that is not surprising for annual data (Table 2). The cointegration test is carried out assuming an intercept in the cointegrating equation.

**TABLE 2**  
Optimum Lag Length through SBC

Lag Length	1	2	3	4
<b>SBC (IND)</b>	-1.758596*	-1.542222	-1.310119	-1.030341
<b>SBC (PAK)</b>	0.272462*	0.584783	0.770143	0.923150
<b>SBC (SRI)</b>	-0.553793*	-0.212490	0.034885	0.220627

Cointegration relationship among real exports, foreign economic activity, relative prices and volatility of exchange rate has been investigated by using the Johansen technique. Table 3 reports our cointegration test results for India, Pakistan and Sri Lanka respectively. Both trace statistics ( $\lambda_{trace}$ ) and maximal eigenvalue ( $\lambda_{max}$ ) statistics indicate that there is at least one cointegrating vector among real exports, foreign economic activity, relative prices and volatility of exchange rate. Therefore, there is a long run equilibrium relationship among all these variables in these three countries. The cointegrating vectors, which are normalized with respect to the real exports, together with their respective  $t$ -values, are given

in Table 4. The results of this normalization yield estimates of the long run elasticities. The signs of the terms are as expected. The relative price term, in other words price elasticity term has a negative sign as expected. The relative price term ranges from 1.153 to 1.484. The foreign activity term is positively related to the real exports and the coefficient of foreign activity term ranges from 2.004 to 2.724. The sign of the exchange rate volatility term for all the three countries is also negative and ranges from 0.994 to 1.463. The negative sign for the volatility term indicates that if volatility (uncertainty) in exchange rate increases, risk-averse producers will favour domestic trade to international trade. Of special interest from these findings is the fact that exchange rate volatility has negative impacts on exports of India, Pakistan and Sri Lanka. This result is in line with findings of *Arize et al. (2000)* and *Baum et al. (2001)* which demonstrate that the negative impact of exchange rate volatility is more clearly observed in less developed countries.

**TABLE 3**  
Cointegration Test Based on Johansen's Maximum Likelihood Method

Country	Trace Statistic				Maximum Eigenvalue Statistic				
	H <sub>0</sub> : H <sub>1</sub> :	r=0 r=1	r=1 r=2	r=2 r=3	r=3 r=4	r=0 r > 0	r≤1 r > 1	r≤2 r > 2	r≤3 r > 3
India	Statistic	68.773***	34.368	15.589	4.174	31.404**	21.779	11.415	4.174
	p-value	0.0014	0.119	0.194	0.387	0.021	0.259	0.223	0.387
Pakistan	Statistic	70.497***	32.824	9.132	3.649	35.673***	20.692	5.483	3.649
	p-value	0.0009	0.145	0.324	0.467	0.005	0.116	0.343	0.467
Sri Lanka	Statistic	95.181***	32.088	8.979	4.147	60.093***	20.109	4.832	4.147
	p-value	0.000	0.151	0.278	0.391	0.000	0.114	0.322	0.391

\*\* and \*\*\* denote rejection of the null hypothesis at the 5 percent and 1 percent significance level respectively.

**TABLE 4**  
Normalized Cointegrating Vector

India	$RX_t = 2.983 -$ (3.849***)	$1.153 RP_t -$ (3.167***)	$1.036 VOL_t +$ (3.854***)	$2.724 Y_t$ (4.466***)
Pakistan	$RX_t = 4.557 -$ (2.509**)	$1.226 RP_t -$ (2.358**)	$0.944 VOL_t +$ (11.801***)	$2.343 Y_t$ (9.816***)
Sri Lanka	$RX_t = 8.839 -$ (4.623***)	$1.484 RP_t -$ (4.088***)	$1.463 VOL_t +$ (3.051***)	$2.01 Y_t$ (7.689***)

Note: The numbers in parentheses beneath the estimated coefficients are t-statistics. \*\* and \*\*\* indicate rejection of the null hypothesis at the 5 % and 1% significance levels respectively.

Since the cointegration tests in the previous section detected one long run equilibrium relationship for each of the export equation, the vector error correction models (VECMs) have been estimated to see stability of the long run equilibrium relationship. Table 5 lists the summary results from the VECMs. We should first note that the one-lagged error-correction term, ECT(-1) appears with a statistically significant coefficient and displays the appropriate (negative) sign in the equations of RX for all the three countries.. The coefficients of the error correction term indicate that the speed of adjustment for Pakistan and Sri Lanka is relatively slow. It means that the adjustment of real exports to any change in the regressors will take a long time to return to the equilibrium. Thus market forces in the export market restore equilibrium slowly in these countries. However, the speed of adjustment in India is relatively high which shows relatively effective role of market forces in the export market for restoring equilibrium a bit speedily. Overall, these findings support the validity of an equilibrium relationship among the variables in the cointegrating equation for all the three countries.

**TABLE 5**  
Summary Results from VECMs

	India	Pakistan	Sri Lanka
	ECT(-1)	ECT(-1)	ECT(-1)
$\Delta(\text{RX})$	-0.616 (-3.158***)	-0.317 (-2.135**)	-0.258 (-1.897*)
$\Delta(\text{RP})$	0.009 (1.161)	0.0023 (0.349)	-0.014 (-0.242)
$\Delta(\text{VOL})$	0.009 (5.448***)	0.0001 (0.654)	0.001 (0.071)
$\Delta(\text{Y})$	-10.6270 (-0.048)	3.342 (1.538)	2.769 (1.524)

*Note: The numbers in parentheses beneath the estimated coefficients are t-statistics. \*, \*\* and \*\*\* denote rejection of the null hypothesis at the 10%, 5% and 1% significance levels respectively.*

## 5. Conclusion and Policy Implications:

The impact of exchange rate volatility on the volume of international trade has been studied intensively since the late 1970's when the exchange rate moved from fixed to flexible exchange rate, means facing a volatile real exchange



rate. The theory says that higher exchange rate volatility will reduce trade by creating uncertainty about future profit from export trade. The aim of this paper is to provide a contribution to the empirical debate on the relationship between exchange rate volatility and exports performance for three South Asian countries India, Pakistan and Sri Lanka. The research exercise has been done using cointegration and vector error-correction model techniques for annual data over the period 1970 to 2007. The results from the cointegration analysis show that there exists a long run equilibrium relationship among real exports demand, relative export prices, exchange rate volatility and foreign economic activity. While the results under the VECMs confirm the stability of this equilibrium relationship among these variables. The overall findings indicate that exchange rate volatility has a negative impact on the exports of India, Pakistan and Sri Lanka. The empirical results derived in this paper are consistent with findings of the studies suggesting that exchange rate volatility in developing economies has a significant negative impact on the export flows to the world market.

The empirical findings of the study point out towards the following policy implications for India, Pakistan and Sri Lanka in order to further improve their exports performance. *First*, for improving the export earnings, these countries should adopt policies with the aim to maintain a stable competitive real exchange rate. In this direction, need is to establish a transparent exchange rate system under which the stability of the real exchange rate is achieved and maintained, and 'getting the exchange rate right' should be the essential part of the overall trade and economic growth strategy. *Secondly*, the finding that foreign economic and relative prices have significant effects on real exports implies that export growth could be driven by factors, which are beyond the control of local policy makers. This implies that external developments are important in influencing export performance.

Our empirical analysis is based on the assumption of a linear relationship among the variables of interest. Possible nonlinear nature of causal links between exports and exchange rate uncertainty may very well be the case. We leave this issue for future work.

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